



# Importance Sampling

## Estimation Theory Project 2

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# Motivation

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- Consider finding the expectation of a function of a random variable

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- Why are we interested in calculating expectation?
  - Many quantities of interest may be cast as expectation





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- *Probabilities:*

$$\mathbb{P}(Y \in A) = \mathbb{E} [I_{\{A\}}(Y)]$$

$$\text{where } I_{\{A\}}(Y) = \begin{cases} 1, & \text{if } Y \in A \\ 0, & \text{if } Y \notin A \end{cases}$$

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- *Integrals:*

$$\int_a^b q(x) dx$$

The above integral can be computed as

$$\begin{aligned} \int_a^b q(x) dx &= (b - a) \int_a^b q(x) \frac{1}{b - a} dx \\ &= (b - a) \int_a^b q(x) P_U(x) dx \\ &= (b - a) \mathbb{E}[q(U)] \end{aligned}$$

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- Remedy?
  - We can compute the expectation by simple Monte-Carlo sampling

# MC Sampling

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# Pitfall of Monte-Carlo Sampling

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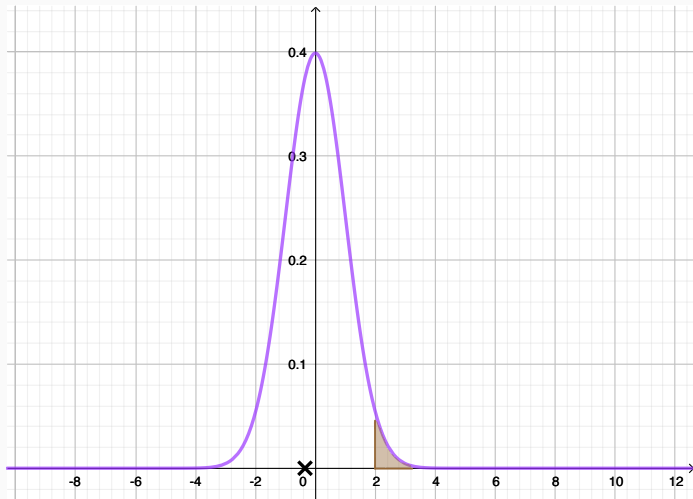


## Pitfall of Monte-Carlo Sampling

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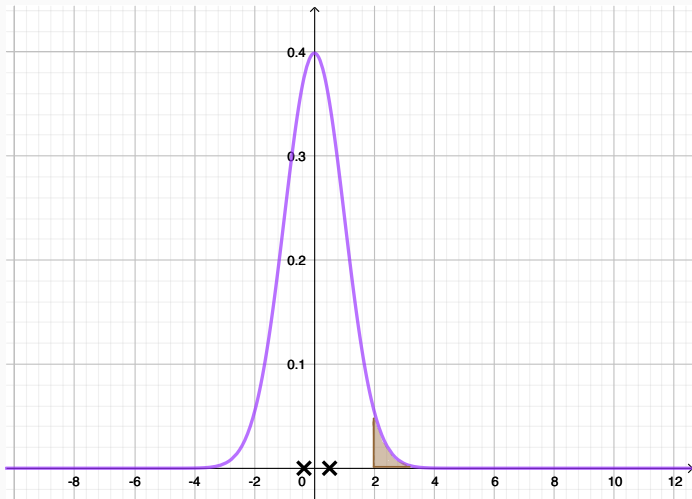
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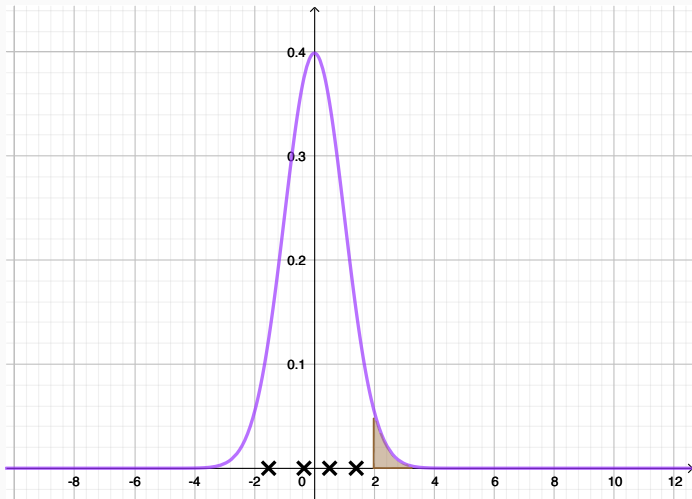
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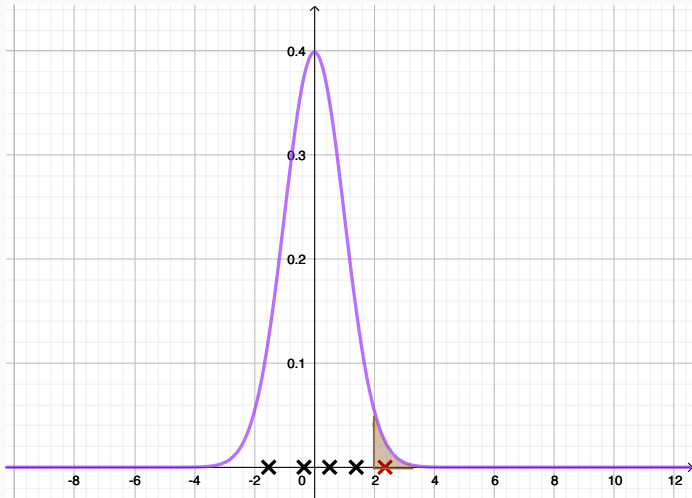
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Importance Sampling!

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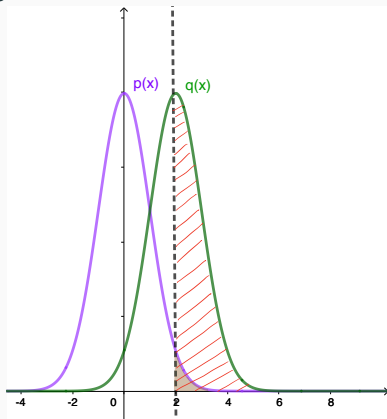
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- Key idea: Sample from a different distribution  $q$  that has “heavy” tail over our region of interest

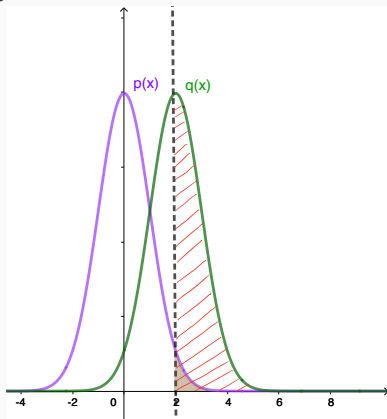
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- $p$  - true distribution,  $q$  - proposal distribution





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- We can rewrite the above expression in terms of our proposal distribution  $q$

$$\mu = \int_{x \in \mathcal{X}} \frac{f(x)p(x)}{q(x)} q(x)dx = \mathbb{E}_q \left[ \frac{f(X)p(X)}{q(X)} \right]$$

$X \stackrel{i.i.d}{\sim} q$

# Requirements for Proposal Distribution

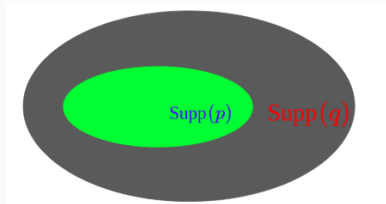
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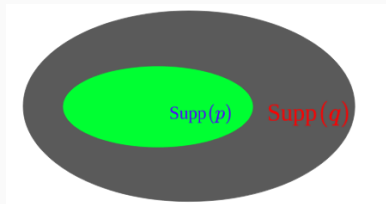


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- $q(x) > 0$  whenever  $f(x)p(x) \neq 0$ 
  - $q(x) > 0$  whenever  $p(x) > 0$



- $\text{Supp}(p) = \{x \in \mathcal{X} : p(x) > 0\}$
- Let  $\mathcal{D}$  denote the support of  $p$  and  $\mathcal{Q}$  denote the support of  $q$

# Importance Sampling Estimate

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• IS estimate:  $\hat{\mu}_q = \frac{1}{N} \sum_{i=1}^N \frac{f(X_i)p(X_i)}{q(X_i)} \quad X_i \stackrel{i.i.d}{\sim} q$

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- IS estimate is unbiased

$$\begin{aligned}\mathbb{E}[\hat{\mu}_q] &= \frac{1}{N} \sum_{i=1}^N \mathbb{E}_q \left[ \frac{f(X_i)p(X_i)}{q(X_i)} \right] \\ &= \mathbb{E}_q \left[ \frac{f(X_1)p(X_1)}{q(X_1)} \right] \\ &= \int_{\mathcal{Q}} \frac{f(x)p(x)}{q(x)} q(x) dx \\ &= \int_{\mathcal{Q}} f(x)p(x) dx\end{aligned}$$

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$$\begin{aligned}\mathbb{E}[\hat{\mu}_q] &\stackrel{(*)}{=} \int_{\mathcal{D}} f(x)p(x) dx + \int_{\mathcal{Q} \cap \mathcal{D}^c} f(x)p(x) dx - \int_{\mathcal{D} \cap \mathcal{Q}^c} f(x)p(x) dx \\ &= \int_{\mathcal{D}} f(x)p(x) dx \\ \implies \mathbb{E}[\hat{\mu}_q] &= \mu\end{aligned}$$

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(\*) follows from the fact that

1.  $\mathcal{Q} \cup (\mathcal{D} \cap \mathcal{Q}^c) = \mathcal{D} \cup (\mathcal{Q} \cap \mathcal{D}^c)$
2.  $\mathcal{Q}$  and  $\mathcal{D} \cap \mathcal{Q}^c$  are disjoint sets, so do  $\mathcal{D}$  and  $\mathcal{Q} \cap \mathcal{D}^c$





## Variance of IS Estimate

$$\cdot \text{Var}_q(\hat{\mu}_q) = \frac{\sigma_q^2}{N} \text{ where}$$

$$\begin{aligned}\sigma_q^2 &= \int_{\mathcal{Q}} \frac{(f(x)p(x))^2}{q(x)} dx - \mu^2 \\ &= \int_{\mathcal{Q}} \frac{(f(x)p(x) - \mu q(x))^2}{q(x)} dx\end{aligned}$$

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- **Note:** Since, for IS,  $\text{supp}(q) \supset \text{supp}(p)$ ,  $\mathcal{Q}$  can be replaced by  $\mathcal{D}$  in the above integral
- since  $\sigma_q^2$  depends on the choice of  $q$ , we can get “optimal”  $q$  that reduces the variance

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Proof:

$$\begin{aligned}\mu^2 + \sigma_{q^*}^2 &= \int_{\mathcal{Q}} \frac{(f(x)p(x))^2}{q^*(x)} dx \\ &= \int_{\mathcal{Q}} \frac{(f(x)p(x))^2}{\frac{|f(x)| p(x)}{\mathbb{E}_p (|f(x)|)}} dx = (\mathbb{E}_p (|f(x)|))^2 = \left( \mathbb{E}_q \left[ \frac{|f(X)| p(X)}{q(X)} \right] \right)^2 \\ &= \left( \int_{\mathcal{Q}} \frac{|f(x)| p(x)}{q(x)} q(x) dx \right)^2 \\ &\stackrel{(*)}{\leq} \int_{\mathcal{Q}} \frac{f^2(x)p^2(x)}{q^2(x)} q(x) dx \underbrace{\int_{\mathcal{Q}} q(x) dx}_{=1} = \mu^2 + \sigma_q^2\end{aligned}$$

$$\implies \sigma_{q^*}^2 \leq \sigma_q^2$$

(\*) follows from Cauchy-Schwartz inequality

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- The appearance of  $q$  in the denominator of  $w$  means that light-tailed  $q$  are dangerous
- so,  $q$  should have tails at least as heavy as  $p$  does

# Simulation Study

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- For comparing MC and IS, we consider the toy problem of calculating the tail of the normal distribution  $\mathcal{N}(0, 1)$

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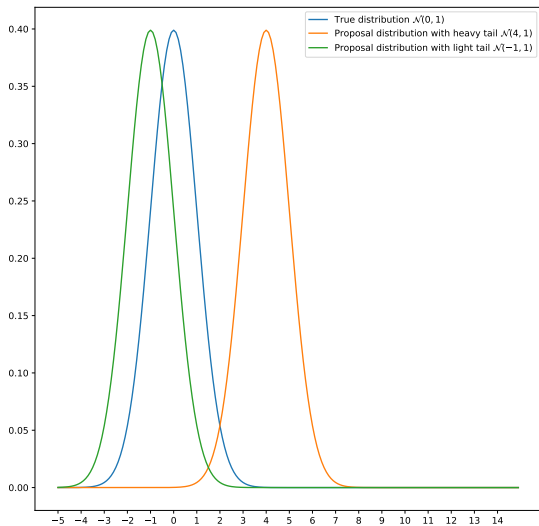
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- We consider two proposal distributions  $q_1$  and  $q_2$  with heavy tail and light tail over  $(3, \infty)$  respectively
  - This is to show how the choice of proposal distribution  $q$  helps or hurts our IS estimate of the tail probability

# Proposal Distributions and True Distribution

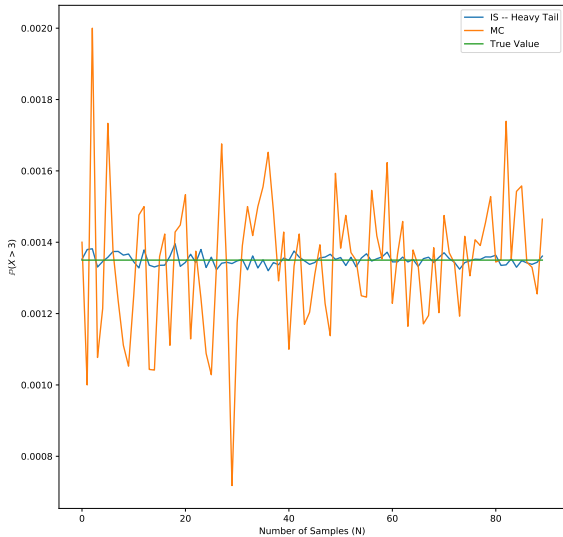


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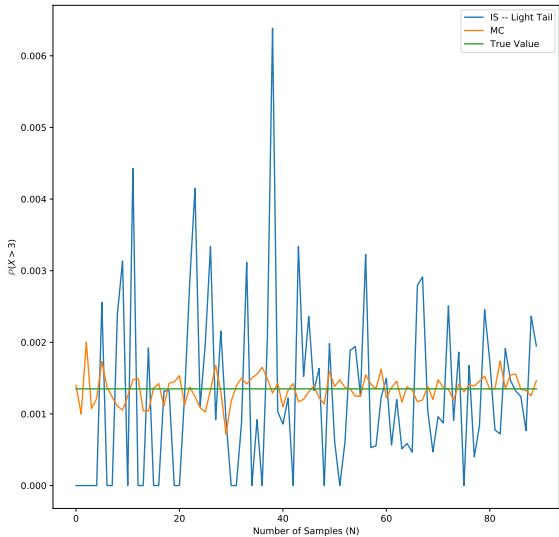
## Results - Using Heavy Tail Proposal Distribution

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## Results - Using Light Tail Proposal Distribution

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# Conclusion

- Monte-Carlo simulation may lead to “poor” approximation i.e, huge variance, in some problems of interest like calculating the tail of a distribution
- This motivated us to look for an alternative sampling technique called *Importance Sampling* which allows us to sample values from the tail using the so called proposal distribution
- We finally conducted a simple toy experiment to see the advantage of using IS. We have also demonstrated how could IS possibly fail

*Thank You*